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Basic Models

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Overview

- Graphical models.
- The *n*-gram model.
- Hidden Markov model.
- Linear classification models.
- Conditional random fields.
- PCFGs
- Forward- and backpropagation in neural networks.
- Recurrent neural networks.
- LSTM networks.

Graphical Models

- Model of the joint distribution $p(X_1, ..., X_N)$ of a set of random variables.
- Direct dependence and independence of variables is represented in the graphical structure.
- Inference: Model allows to infer
 - All marginal probabilities $p(X_{i_1}, ..., X_{i_m})$,
 - All conditional probabilities $p(X_{i_1}, ..., X_{i_k} | X_{i_{k+1}}, ..., X_{i_m})$.
- Random variables can represent acoustic signals, letters, words, parts of speech, semantic labels, ...

Directed Graphical Models

- Directed graphical model is a graph over
 - Nodes: set of random variables $\{X_1, ..., X_N\}$.
 - With edges $\subset \{X_1, ..., X_N\}^2$ that contain no cycles.
 - Edges are written $X_i \rightarrow X_i$.
- Parents $pa(X_i) = \{X_j | X_j \rightarrow X_i\}$ are the nodes which X_i is directly dependent on.
- The directed model represents a joint distribution
 - $P(X_1, ..., X_N) = \prod_{i=1}^N p(X_i | pa(X_i)).$

Directed Graphical Models

- Why does the graph have to be acyclic?
- It the graph is acyclic then there is an ordering $i_1, ..., i_N$ of the nodes such that
 - For all i_j and for all $X_{i_k} \in pa\left(X_{i_j}\right) : k < j$.
 - That is, the parents come before their children in the odering.
- For such an ordering, we can factorize

•
$$P(X_1, ..., X_N) = \prod_{i=1}^N p(X_i | pa(X_i)).$$

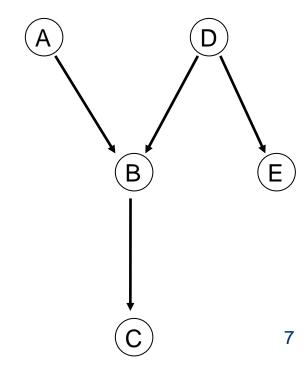
Before X_i in the ordering

Directed Graphical Models: Independence

- The graph structure of a graphical model implies (conditional) independencies between random variables.
 - $\bullet \neg (X_i \to X_j) \text{ implies } p(X_j | pa(X_j), X_i) = p(X_j | pa(X_j)).$
- Independences make inference easier!
 - Depending on the structure of the model, polynomial instead of exponential complexity.
- Dependences and independences represent domain knowledge and modeling assumptions.

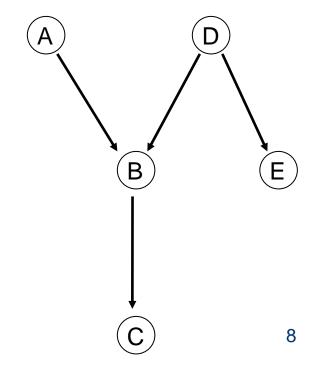
Directed Graphical Models: Example

Ordering:



Directed Graphical Models: Example

- Ordering: A, D, B, E, C.
- P(A, B, C, D, E) = P(A)P(D)P(B|A, D)P(E|D)P(C|B).

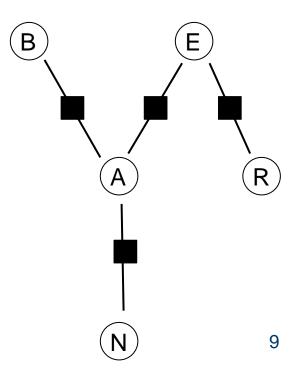


Undirected Graphical Models

• Undirected graphical model over $X_1, ..., X_N$:

$$P(X_1, \dots, X_N) = \frac{1}{Z} \prod_{i=1}^k \Psi_i$$

- Represented by a factor graph
- Solid nodes are factors.
- Each factor is joint distribution of its connected nodes.

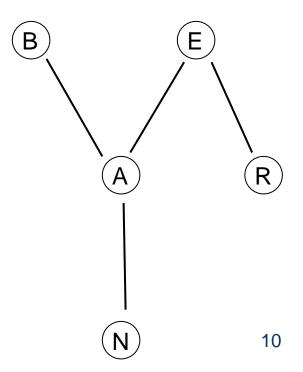


Undirected Graphical Models

• Undirected graphical model over $X_1, ..., X_N$:

$$P(X_1, \dots, X_N) = \frac{1}{Z} \prod_{i=1}^k \Psi_i$$

- Factor graph represents factorization
- Markov field: connect nodes that occur in a joint factor.
- Markov field reflects conditional independence.



Undirected Graphical Models

• Undirected graphical model over X_1, \dots, X_N :

$$P(X_1, \dots, X_N) = \frac{1}{Z} \prod_{j=1}^k \Psi_j$$

• With $\Psi_i = e^{\psi_j}$:

$$P(X_1, ..., X_N) = \frac{1}{Z} \exp \left\{ \sum_{j=1}^k \psi_j \right\}$$

Inference in Graphical Models

- Problem setting for inference.
- Given observations for some variables X_{i_1}, \dots, X_{i_m} ,
- Infer distribution of query variable X_q :
 - $p(X_q | x_{i_1}, ..., x_{i_m}).$
- Cannot immediately calculate value because the values of the unobserved parents are unknown.
 - Sum over all values (summation rule).
- Notation: $\{X_1, \dots, X_N\} = \{\underbrace{X_q}_{\text{query}}, \underbrace{X_{i_1}, \dots, X_{i_m}}_{\text{observed}}, \underbrace{X_{j_1}, \dots, X_{j_k}}_{\text{unobserved}}\}$

Graphical Models: Inference

- Notation: $\{X_1, \dots, X_N\} = \{\underbrace{X_q}_{\text{query}}, \underbrace{X_{i_1}, \dots, X_{i_m}}_{\text{observed}}, \underbrace{X_{j_1}, \dots, X_{j_k}}_{\text{unobserved}}\}$
- Inference:

$$p(X_{q}|x_{i_{1}},...,x_{i_{m}}) = \frac{p(X_{q},x_{i_{1}},...,x_{i_{m}})}{p(x_{i_{1}},...,x_{i_{m}})}$$

$$= \frac{1}{p(x_{i_{1}},...,x_{i_{m}})} \sum_{x_{j_{1}}} ... \sum_{x_{j_{k}}} p(X_{q},x,...,x_{i_{m}},x_{j_{1}},...,x_{j_{k}})$$

Number of summands is exponential in the number of unobserved variables.

Graphical Models: Inference

Inference:

$$p(X_q | x_{i_1}, \dots, x_{i_m})$$

$$= \frac{1}{p(x_{i_1}, \dots, x_{i_m})} \sum_{x_{j_1}} \dots \sum_{x_{j_k}} p(X_q, x, \dots, x_{i_m}, x_{j_1}, \dots, x_{j_k})$$

- Exact inference:
 - Message passing algorithm.
 - For general graph structures intractable.
 - If the model has sequential structure: quadratic (Viterbi-/forward-backward algorithm).
 - If the model has tree structure: cubic.

Graphical Models: Inference

Inference:

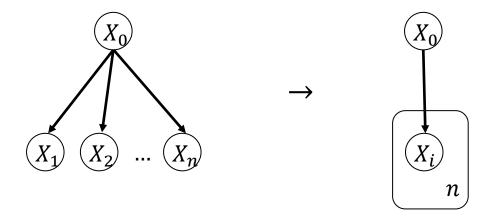
$$p(X_q | x_{i_1}, ..., x_{i_m})$$

$$= \frac{1}{p(x_{i_1}, ..., x_{i_m})} \sum_{x_{j_1}} ... \sum_{x_{j_k}} p(X_q, x, ..., x_{i_m}, x_{j_1}, ..., x_{j_k})$$

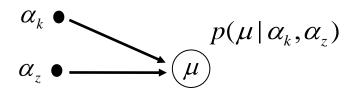
- Approximate inference for general graph structures
 - Loopy belief propagation,
 - Variational methods.

Plate Notation

Shorthand notation for "loops":



Notation for parameters of distributions:



Overview

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- LSTM networks.

- Basic tool for language modeling.
- Unigram model (n = 1)
- Based on the Markov assumption of order 0:

•
$$p(X_t|X_{t-1},...,X_1) = p(X_t)$$

$$X_1$$

 X_2

 X_3

 X_4

 (X_5)

 (X_6)

...

 (X_T)

$$p(X_1, ..., X_T) = p(X_1)p(X_2)p(X_3)p(X_4)p(X_5)p(X_6) ... p(X_T)$$

$$= \prod_{i=1}^{T} p(X_i)$$

- Basic tool for language modeling.
- Bigram model (n = 2)
- Based on the Markov assumption of order 1:

•
$$p(X_t|X_{t-1},...,X_1) = p(X_t|X_{t-1})$$

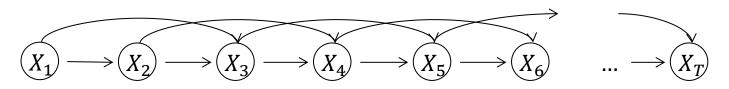
$$(X_1) \longrightarrow (X_2) \longrightarrow (X_3) \longrightarrow (X_4) \longrightarrow (X_5) \longrightarrow (X_6) \qquad \dots \longrightarrow (X_T)$$

$$p(X_1, ..., X_T) = p(X_1)p(X_2|X_1)p(X_3|X_2) ... p(X_T|X_{T-1})$$

$$= P(X_1) \prod_{t=2}^{T} p(X_t|X_{t-1})$$

- Basic tool for language modeling.
- Trigram model (n = 3)
- Based on the Markov assumption of order 3:

•
$$p(X_t|X_{t-1},...,X_1) = p(X_t|X_{t-1},X_{t-2})$$



$$p(X_1, ..., X_T)$$

$$= p(X_1)p(X_2|X_1)p(X_3|X_2, X_1) ... p(X_T|X_{T-1}, X_{T-2})$$

$$= P(X_1)p(X_2|X_1) \prod_{t=3}^{T} p(X_t|X_{t-1}, X_{t-2})$$

- Basic tool for language modeling.
- Based on the Markov assumption of order n − 1:

$$p(X_t | X_{t-1}, ..., X_1) = p(X_t | X_{t-1}, ..., X_{t-n+1})$$

n-gram model:

$$\begin{aligned} p(X_1, \dots, X_T) \\ &= p(X_1) \dots p(X_T | X_{T-1}, \dots, X_{T-n+1}) \\ &= P(X_1) \dots p(X_{n-1} | X_{n-2}, \dots, X_1) \prod_{t=n}^T p(X_t | X_{t-1}, \dots, X_{t-n+1}) \\ &\text{Categorical distributions} \end{aligned}$$

- Basic tool for language modeling.
- Based on the Markov assumption of order n − 1:

$$p(X_t | X_{t-1}, ..., X_1) = p(X_t | X_{t-1}, ..., X_{t-n+1})$$

n-gram model:

$$\begin{aligned} p(X_1, ..., X_T) &= p(X_1) ... p(X_T | X_{T-1}, ..., X_{T-n+1}) \\ &= P(X_1) ... p(X_{n-1} | X_{n-2}, ..., X_1) \prod_{t=n}^{T} p(X_t | X_{t-1}, ..., X_{t-n+1}) \\ &\text{Parameters of the } n\text{-gram model} \end{aligned}$$

Parameters of the *n*-Gram Model

 Parameters: values that are required to infer the likelihood of an observation in the model.

•
$$p(x_n|x_{n-1},...,x_1) = \frac{p(x_1,...,x_n)}{p(x_1,...,x_{n-1})}$$

- For each m with $1 \le m \le n$ and each combination of values $x_1, ..., x_m$:
 - $p(x_1, ..., x_m)$ has to be known.
 - Written as $\theta_{x_1,...,x_m}$.
- We will infer these parameters from data (e.g., from a text corpus).

The *n*-Gram Model: Parameter Inference

- Given data from which we determine:
 - For all m with $1 \le m \le n$ and each combination of values $x_1, ..., x_m$ we observe the number of occurances $N_{x_1,...,x_m}$.
 - For all m with $1 \le m \le n$ we observe the total number N_m of observations.
- What is the relationship between the true, unknown parameters $\theta_{x_1,...,x_m}$ and the observed $N_{x_1,...,x_m}$?

The *n*-Gram Model: Parameter Inference

- What is the relationship between the true, unknown parameters $\theta_{x_1,...,x_m}$ and the observed $N_{x_1,...,x_m}$?
 - Each $N_{x_1,...,x_m}$ is a random variable.
 - \bullet N_m random experiments,
 - Each $(x_1, ..., x_m) \in Y^m$ is a possible outcome that occurs with probability $\theta_{x_1,...,x_m}$.
 - Let $\{y_1, ..., y_k\} = Y^m$ be all the possible outcomes
 - What type of distribution is $p(N_{y_1}, ..., N_{y_k} | \theta_{y_1}, ..., \theta_{y_k})$?

The *n*-Gram Model: Parameter Inference

- What is the relationship between the true, unknown parameters $\theta_{x_1,...,x_m}$ and the observed $N_{x_1,...,x_m}$?
 - Each $N_{x_1,...,x_m}$ is a random variable.
 - \bullet N_m random experiments,
 - Each $(x_1, ..., x_m) \in Y^m$ is a possible outcome that occurs with probability $\theta_{x_1,...,x_m}$.
 - Let $\{y_1, ..., y_k\} = Y^m$ be all the possible outcomes
- $p(N_{y_1}, ..., N_{y_k} | \theta_{y_1}, ..., \theta_{y_k})$ is a multinomial distribution.

•
$$p(N_{y_1}, ..., N_{y_k} | \theta_{y_1}, ..., \theta_{y_k}) = \frac{N_m!}{N_{y_1}!...N_{y_k}!} \theta_{y_1}^{N_{y_1}} \cdot ... \cdot \theta_{y_k}^{N_{y_k}}$$

with $N_m = \sum_j N_{y_j}$

The *n*-Gram Model: ML Parameters

Likelihood of training data:

$$p(N_{y_1}, ..., N_{y_k} | \theta_{y_1}, ..., \theta_{y_k}) = M[\theta_{y_1}, ..., \theta_{y_k}](N_{y_1}, ..., N_{y_k})$$

$$= \frac{N_m!}{N_{y_1}! ... N_{y_k}!} \theta_{y_1}^{N_{y_1}} \cdot ... \cdot \theta_{y_k}^{N_{y_k}}$$

Maximum-likelihood estimate:

$$\bullet \arg \max_{\theta_{y_1}, \dots, \theta_{y_k}} M[\theta_{y_1}, \dots, \theta_{y_k}](N_{y_1}, \dots, N_{y_k})$$

•
$$\theta_{y_i}^{\text{ML}} = \frac{N_{y_i}}{N_m} = \frac{\text{Number of occurrences of } y_i}{\text{Number of observed } n\text{-gram combinations}}$$

 Maximum-likelihood parameters generally not robust, unregularized estimates.

The *n*-Gram Model: MAP Parameters

Posterior of training data:

$$p(\theta_{y_1}, ..., \theta_{y_k} | N_{y_1}, ..., N_{y_k})$$

$$= \frac{1}{p(N_{y_1}, ..., N_{y_k})} p(N_{y_1}, ..., N_{y_k} | \theta_{y_1}, ..., \theta_{y_k}) p(\theta_{y_1}, ..., \theta_{y_k})$$

- Likelihood $p(N_{y_1}, ..., N_{y_k} | \theta_{y_1}, ..., \theta_{y_k})$ is a multinomial distribution.
- If prior $p(\theta_{y_1}, ..., \theta_{y_k})$ follows a Dirichlet distribution, then posterior $p(\theta_{y_1}, ..., \theta_{y_k} | N_{y_1}, ..., N_{y_k})$ follows a Dirichlet distribution as well.
 - Dirichlet distribution is the conjugate of the multinomial distribution.

The *n*-Gram Model: MAP Parameters

Posterior of training data:

$$p(\theta_{y_{1}}, ..., \theta_{y_{k}} | N_{y_{1}}, ..., N_{y_{k}})$$

$$= \frac{1}{p(N_{y_{1}}, ..., N_{y_{k}})} M[\theta_{y_{1}}, ..., \theta_{y_{k}}](N_{y_{1}}, ..., N_{y_{k}}) D[\alpha_{y_{1}}, ..., \alpha_{y_{k}}] p(\theta_{y_{1}}, ..., \theta_{y_{k}})$$

$$= \frac{1}{p(N_{y_{1}}, ..., N_{y_{k}})} D[\alpha_{y_{1}} + \theta_{y_{1}}, ..., \alpha_{y_{k}} + \theta_{y_{k}}](N_{y_{1}}, ..., N_{y_{k}})$$

Maximum-posterior estimate:

$$\bullet \quad \arg\max_{\theta_{y_1},\dots,\theta_{y_k}} D[\alpha_{y_1}+\theta_{y_1},\dots,\alpha_{y_k}+\theta_{y_k}](N_{y_1},\dots,N_{y_k})$$

$$\theta_{y_i}^{\text{MAP}} = \frac{N_{y_i} + \alpha_{y_i}}{N_m + \sum_{y} \alpha_y} =$$
Number of occurrences of $y_i + \alpha_{y_i}$

Number of observed n-gram combinations+ $\sum_{y} \alpha_{y}$

 This form of regularization is also called Laplace smoothing.

Overview

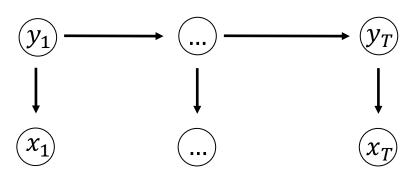
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The Hidden Markov Model

- Directed, generative model.
- Joint probability of input and output:

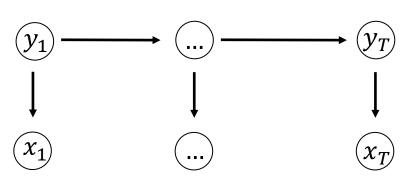
$$P(\mathbf{x}, \mathbf{y}|\mathbf{\theta}) = P(x_1, \dots, x_T, y_1, \dots, y_T|\mathbf{\theta})$$
$$= \prod_{t=1}^{T} P(y_t|y_{t-1}, \mathbf{\theta}) P(x_t|y_t, \mathbf{\theta})$$

 Generative model: parameters optimized to maximize joint likelihood of input and output.



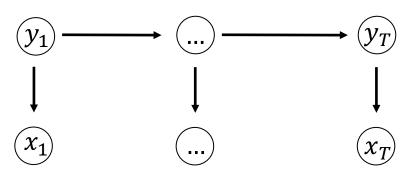
The Hidden Markov Model

- Model parameters in vector $\mathbf{\theta} = (\boldsymbol{\pi}, \mathbf{a}, \mathbf{b})$:
 - Starting state probabilities $\pi_i = P(y_1 = i | \theta)$
 - Transition probabilities $a_{ij} = P(y_{t+1} = j | y_t = i, \theta)$
 - Observation probabilities $b_i(o) = P(x_t = o | y_t = i, \theta)$
- Markov assumptions:
 - $P(y_{t+1}|y_1,...,y_t,\mathbf{\theta}) = P(y_{t+1}|y_t,\mathbf{\theta})$
 - $\bullet P(x_t|y_1,...,y_T,\mathbf{\theta}) = P(x_t|y_t,\mathbf{\theta})$



The Hidden Markov Model

- Prediction, inference: most likely output given input $\underset{\text{argmax}_{y_1,...,y_T}}{\text{ergmax}_{y_1,...,y_T}} P(y_1,...,y_T | x_1,...,x_T, \mathbf{\theta})$
- Naive maximization over all combinations of labels; exponentially many in T.
- But can be solved in O(T) with dynamic programming \rightarrow Viterbi algorithm.



HMM: Inference Tasks

- Find most likely state sequence given output:
 - $\operatorname{argmax}_{y_1, \dots, y_T} P(y_1, \dots, y_T | x_1, \dots, x_T, \boldsymbol{\theta})$
 - ullet Solved in O(T) by the Viterbi algorithm.
- Likelihood of an observation sequence:
 - $P(x_1, ..., x_T | \mathbf{\theta})$
 - \bullet Solved in O(T) by the forward-backward algorithm.
- Find most likely state at t given output sequence:
 - $\operatorname{argmax}_{y_t} P(y_t | x_1, ..., x_T, \boldsymbol{\theta})$
 - * Solved in O(T) by the forward-backward algorithm. \downarrow

Viterbi Algorithm

Definition:

$$\delta_t(i) = \max_{y_1, \dots, y_{t-1}} P(y_1, \dots, y_{t-1}, y_t = i, x_1, \dots, x_t | \mathbf{\theta})$$

- It follows that $\sum_i \delta_T(i) = P(\mathbf{x}, \mathbf{y} | \mathbf{\theta})$
- Since x is constant, maximizing $\sum_i \delta_T(i)$ over y maximizes $P(y|x, \theta)$.

Viterbi Algorithm

Definition:

$$\delta_t(i) = \max_{y_1, \dots, y_{t-1}} P(y_1, \dots, y_{t-1}, y_t = i, x_1, \dots, x_t | \mathbf{\theta})$$

Theorem: can be calculated recursively

$$\delta_1(i) = p(y_1 = i|\mathbf{\theta})P(x_1|y_1 = i,\mathbf{\theta})$$

$$\delta_{t+1}(j) = (\max_i \delta_t(i)a_{ij})b_j(x_{t+1})$$

Proof:

$$\begin{split} &\delta_{t+1}(j) = \max_{y_1, \dots, y_t} P(y_1, \dots, y_t, y_{t+1} = j, x_1, \dots, x_{t+1} | \mathbf{\theta}) \\ &= \max_{y_1, \dots, y_t} P(y_1, \dots, y_t, x_1, \dots, x_t | \mathbf{\theta}) a_{ij} b_j(x_{t+1}) \\ &= \max_{i} \max_{y_1, \dots, y_{t-1}} P(y_1, \dots, y_t = i, x_1, \dots, x_t | \mathbf{\theta}) a_{ij} b_j(x_{t+1}) \\ &= \max_{i} \delta_t(i) a_{ij} b_j(x_{t+1}) \end{split}$$

Viterbi Algorithm

- Initialization:
 - $\bullet \log \delta_1(i) = \log \pi_i b_i(x_1)$
 - $\psi_1(i) = 0$
- For t = 1 ... T 1, for all y':
 - $\bullet \log \delta_{t+1}(j) = \left(\max_{i} \log \delta_{t}(i) + \log a_{ij}\right) + \log b_{j}(x_{t+1})$
 - $\psi_{t+1}(j) = \left(\arg\max_{i} \log \delta_t(i) + a_{ij}\right)$
- Termination: $y_T^* = \arg \max_{y} \delta_T(y)$
- For $t = T 1 \dots 1$
 - $y_t^* = \psi_{t+1}(q_{t+1}^*)$
- Return $y_1^*, ..., y_T^*$.

Forward-Backward Algorithm

- Definitions:
- Relation to inference problems:
 - Likelihood of observation sequence:

$$P(O_1, ..., O_T | \theta) = \sum_i \alpha_T(i).$$

Most likely state given observation sequence:

$$P(q_t = i | O_1, ..., O_T, \theta) = \gamma_t(i).$$

Forward Step

Likelihood of an initial observation sequence:

Theorem:

Forward algorithm:

- lacktriangledow For all i, let $lpha_1(i)=\pi_i b_i(O_1)$
- For $t=1\dots T-1$: let $\alpha_{t+1}(j) = \left(\sum_{i=1}^N \alpha_t(i)a_{ij}\right)b_j(O_{t+1})$

Forward Step: Proof

Proof by induction: base case

•
$$\alpha_1(i) = P(O_1, q_1 = i \mid \lambda)$$

= $P(q_1 = i \mid \lambda)P(O_1 \mid q_1 = i, \lambda) = \pi_i b_i(O_1)$

■ Induction: $t \rightarrow t + 1$

$$\alpha_{t+1}(j) = P(O_1, ..., O_{t+1}, q_{t+1} = j \mid \lambda) = \sum_{i=1}^{N} P(O_1, ..., O_{t+1}, q_t = i, q_{t+1} = j \mid \lambda)$$

$$= \sum_{i=1}^{N} P(O_1, ..., O_t, q_t = i \mid \lambda) P(q_{t+1} = j \mid q_t = i, O_1, ..., O_t, \lambda)$$

$$P(O_{t+1} \mid q_{t+1} = j, q_t = i, O_1, ..., O_t, \lambda)$$

$$= \sum_{i=1}^{N} P(O_1, ..., O_t, q_t = i \mid \lambda) P(q_{t+1} = j \mid q_t = i, \lambda) P(O_{t+1} \mid q_{t+1} = j, \lambda)$$

$$= \left(\sum_{i=1}^{N} \alpha_t(i) a_{ij}\right) b_j(O_{t+1})$$

Backward Step

Likelihood of rest of observation sequence:

Theorem:

- $+ \beta_T(i) = 1$

Backward algorithm:

- For all i, let $\beta_T(i) = 1$
- For $t=T-1\dots 1$, for all i: let $\beta_t(i)=\left(\sum_{j=1}^N a_{ij}\,b_j(O_{t+1})\beta_{t+1}(j)\right)$

Backward Step: Proof

■ Induction step $t + 1 \rightarrow t$

$$\begin{split} &\beta_{t}(i) = P(O_{t+1}, ..., O_{T} \mid q_{t} = i, \lambda) \\ &= \sum_{j} P(O_{t+1}, ..., O_{T}, q_{t+1} = j \mid q_{t} = i, \lambda) \\ &= \sum_{j} P(O_{t+1}, ..., O_{T} \mid q_{t+1} = j, q_{t} = i, \lambda) P(q_{t+1} = j \mid q_{t} = i, \lambda) \\ &= \sum_{j} P(O_{t+2}, ..., O_{T} \mid q_{t+1} = j, q_{t} = i, \lambda) P(q_{t+1} = j \mid q_{t} = i, \lambda) \\ &P(O_{t+1} \mid q_{t+1} = j, q_{t} = i, \lambda) \\ &= \sum_{j} P(O_{t+2}, ..., O_{T} \mid q_{t+1} = j, \lambda) P(q_{t+1} = j \mid q_{t} = i, \lambda) P(O_{t+1} \mid q_{t+1} = j, \lambda) \\ &= \sum_{j} \beta_{t+1}(j) a_{ij} b_{j}(O_{t+1}) \end{split}$$

Forward-Backward Algorithm

Probability of a state at time t:

$$\begin{aligned} & \quad \gamma_{t}(i) = P(q_{t} = i | O_{1}, \dots, O_{T}, \theta) \\ & = \frac{P(q_{t} = i, O_{1}, \dots, O_{T} | \theta)}{P(O_{1}, \dots, O_{T} | \theta)} \\ & = \frac{P(q_{t} = i, O_{1}, \dots, O_{t}, O_{t+1}, O_{T} | \theta)}{P(O_{1}, \dots, O_{T} | \theta)} \\ & = \frac{P(q_{t} = i, O_{1}, \dots, O_{t}) P(O_{t+1}, O_{T} | q_{t} = i, \theta)}{P(O_{1}, \dots, O_{T} | \theta)} \\ & = \frac{\alpha_{t}(i) \beta_{t}(i)}{\sum_{i} \alpha_{T}(i)} \end{aligned}$$

Forward-Backward Algorithm

- Forward pass:
 - lacktriangledow For all i, let $lpha_1(i)=\pi_i b_i(O_1)$
 - For $t = 1 \dots T 1$, for all i:
 - \star Let $\alpha_{t+1}(j) = \left(\sum_{i=1}^{N} \alpha_t(i) a_{ij}\right) b_j(O_{t+1})$
- Backward pass:
 - lacktriangle For all i, let $eta_T(i)=1$
 - For $t = T 1 \dots 1$, for all i:
 - \star Let $\beta_t(i) = \left(\sum_{j=1}^N a_{ij} b_j(O_{t+1}) \beta_{t+1}(j)\right)$
 - \star Let $\gamma_t(i) = \frac{\alpha_t(i)\beta_t(i)}{\sum_i \alpha_T(i)}$

Learning in Hidden Markov Models

- Focus on discrete observations O_t first.
- Hidden states are visible during training:
 - Training data are labeled with states.
 - 2. For instance, in part-of-speech tagging.
- 2. Hidden states are not visible during training:
 - States have to be inferred.
 - For instance, in speech recognition.
 - 3. Training by Baum-Welch algorithm.

HMM Learning with Visible States

• Maximum-likelihood parameters $\theta = (\pi, a, b)$:

•
$$\pi_i = P(y_1 = i | \mathbf{\theta}) = \frac{\text{\# sequences that start in state } i}{\text{\# sequences}}$$

•
$$a_{ij} = P(y_{t+1} = j | y_t = i, \mathbf{\theta}) = \frac{\text{# state transitions from } i \text{ to } j}{\text{# state transitions from } i}$$

•
$$b_i(o) = P(x_t = o | y_t = i, \mathbf{\theta}) = \frac{\text{# of times } x_t = o \text{ in state } i}{\text{# of times in state } i}$$

 Laplace correction (regularization): add constant to numerator, change denominator so that probabilities add to one again.

HMM Learning with Invisible States

- Start with initial, random model
- Iteratively:
 - use forward backward to estimate state probabilities
 - Update model parameters based on estimated states.
- Baum-Welch algorithm is variant of the EM algorithm for the hidden markov model.

Baum-Welch Algorithm

Definition: probability of a transition from state i to state j at time t:

Can be calculated as:

•
$$\xi_t(i,j) = P(q_t = i, q_{t+1} = j | O_1, ..., O_T, \theta)$$

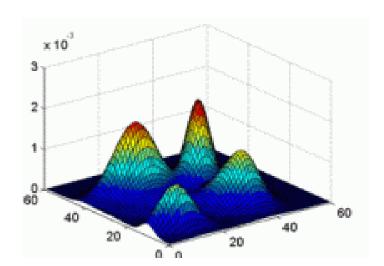
= $\frac{\alpha_t(i) a_{ij} \beta_{t+1}(j)}{\sum_{i=1}^N \alpha_T(i)}$

Baum-Welch Algorithm

- Input: Set of observation sequences $\{(O_1^1, ..., O_T^1), ..., (O_1^m, ..., O_T^m)\}$
- lacktriangleright Initialize heta at random
- Repeat until convergence:
 - Run forward-backward algorithm to infer all $\alpha_t(i)$, $\beta_t(i)$, $\gamma_t(i)$.
 - lacktriangle For all t, i, j: infer $\xi_t(i,j)$
 - For all i: let $\pi_i = \gamma_1(i)$
 - For all i, j: let $a_{ij} = \frac{\sum_{t=1}^{T} \xi_t(i,j)}{\sum_{t=1}^{T} \gamma_t(i)}$
 - For all i, 0: let $b_i(0) = \frac{\sum_{t:O_t=0} \gamma_t(i)}{\sum_{t=1}^T \gamma_t(i)}$

Continuous Hidden Markov Models

- So far, $b_i(O_t)$ has been a multinomial distribution over a set of discrete observations.
- In speech processing, O_t is a vector of continuous attributes.
- Modeling assumption: mixture of Gaussian distributions.
 - $b_i(\mathbf{x}_t) = \sum_{k=1}^{M} c_{ik} N[\mu_{ik}, \Sigma_{ik}](\mathbf{x}_t) = \sum_{k=1}^{M} c_{ik} b_{ik}(\mathbf{x}_t)$



Continuous Hidden Markov Models

Modeling assumption: mixture of Gaussian distributions.

•
$$b_i(\mathbf{x}_t) = \sum_{k=1}^{M} c_{ik} N[\mu_{ik}, \Sigma_{ik}](\mathbf{x}_t) = \sum_{k=1}^{M} c_{ik} b_{ik}(\mathbf{x}_t)$$

- Parameters of multivariate Gaussian distributions $b_{ik}(\mathbf{x}_t)$ have to be estimated from data.
- Mixing coefficients c_{ik} are unknown; estimated using EM algorithm.
- Each iteration of the Baum-Welch algorithm requires an execution of the EM algorithm to estimate mixing coefficients c_{ik} .

The Hidden Markov Model

- Generative sequence model.
- Observed output variable, latent state variable.
- Latent states, state sequences can be inferred by Viterbi-/forward-backward-algorithm.
- States visible in labeled training data? Model parameters estimated by "regularized counting".
- States invisible in labeled training data? Model parameters estimated by Baum-Welch algorithm.
- Continuous observations usually modeled as (mixture of) multivariate Gaussian distributions.

Overview

- Graphical models.
- The n-gram model.
- Hidden Markov model.
- Linear classification models.
- Conditional random fields.
- PCFGs
- Forward- and backpropagation in neural networks.
- Recurrent neural networks.
- LSTM networks.

Classification

- Input: an instance $x \in X$
 - ◆ E.g., X can be a vector space over attributes
 - The Instance is then an assignment of attributes.

•
$$\mathbf{x} = \begin{pmatrix} x_1 \\ \vdots \\ x_m \end{pmatrix}$$
 is a feature vector

- Output: Class $y \in Y$; where Y is a finite set.
 - The class is also referred to as the target attribute
 - y is also referred to as the (class) label

$$\mathbf{x} \rightarrow \mathbf{classifier} \rightarrow y$$

Classification: Example

- Input: Instance $x \in X$
 - X: the set of all possible combinations of regiment of medication



• Output: $y \in Y = \{\text{toxic, ok}\}$



Linear Classification Models

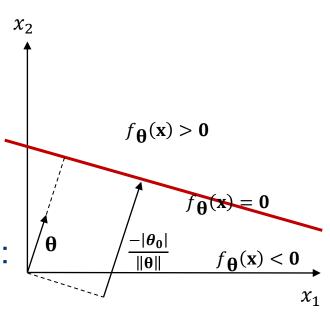
• Hyperplane given by normal vector & displacement: $H_{\mathbf{\theta}} = \{\mathbf{x} | f_{\mathbf{\theta}}(\mathbf{x}) = \mathbf{x}^{\mathrm{T}} \mathbf{\theta} + \theta_0 = 0\}$

- Example: $X = \mathbb{R}^2$
- Decision function:

$$f_{\boldsymbol{\theta}}(\mathbf{x}) = \mathbf{x}^{\mathrm{T}} \mathbf{\theta} + \theta_0$$

■ Binary classifier, $y \in \{+1, -1\}$:

$$y_{\theta}(\mathbf{x}) = \operatorname{sign}(f_{\theta}(\mathbf{x}))$$



Linear Classification Models

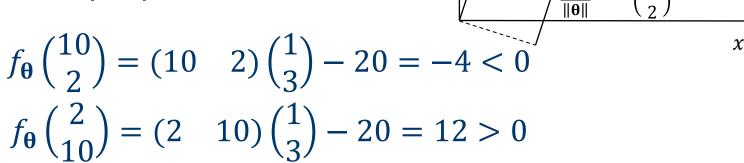
Hyperplane given by normal vector & displacement:

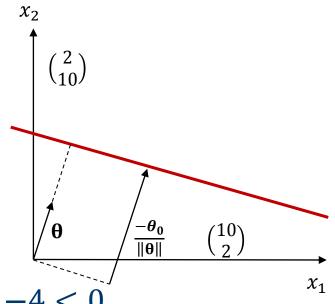
$$H_{\boldsymbol{\theta}} = \{ \mathbf{x} | f_{\boldsymbol{\theta}}(\mathbf{x}) = \mathbf{x}^{\mathrm{T}} \boldsymbol{\theta} + \theta_0 = 0 \}$$

- Example: $X = \mathbb{R}^2$
- Decision function:

$$f_{\mathbf{\theta}}(\mathbf{x}) = \mathbf{x}^{\mathrm{T}} \begin{pmatrix} 1 \\ 3 \end{pmatrix} - 20$$

Example points:





Linear Classification Model

- Offset can "disappear" into parameter vector.
- Example
 - Before: $f_{\theta}(\mathbf{x}) = \begin{pmatrix} x_1 \\ x_2 \end{pmatrix}^{\mathrm{T}} \begin{pmatrix} 1 \\ 3 \end{pmatrix} 20$

• After:
$$f_{\theta}(\mathbf{x}) = \begin{pmatrix} 1 \\ x_1 \\ x_2 \end{pmatrix}^{\mathrm{T}} \begin{pmatrix} -20 \\ 1 \\ 3 \end{pmatrix}$$

- New constant attribute $x_0 = 1$ added to all instances
- Offset θ_0 integrated into θ .

Learning Linear Classifiers

Input to the Learner: Training data T_n .

$$\mathbf{X} = \begin{pmatrix} x_{11} & \cdots & x_{1m} \\ \vdots & \ddots & \vdots \\ x_{n1} & \cdots & x_{nm} \end{pmatrix}$$

$$\mathbf{Y}_{\mathbf{0}} : X \to Y$$

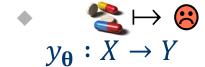
$$\mathbf{Linear classifier:}$$

$$\mathbf{y} = \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix}$$

Training Data:

$$T_n = \{(\mathbf{x}_1, y_1), \dots, (\mathbf{x}_n, y_n)\}$$

Output: a Model



$$y_{\mathbf{\theta}}(\mathbf{x}) = \begin{cases} \mathbf{\Theta} & \text{if } \mathbf{x}^{\mathrm{T}} \mathbf{\theta} \ge 0 \\ & \text{otherwise} \end{cases}$$

Linear classifier with parameter vector $\boldsymbol{\theta}$.

Solve

$$\underset{\boldsymbol{\theta}}{\operatorname{argmin}} \sum_{i=1}^{n} \ell(f_{\boldsymbol{\theta}}(\mathbf{x}_i), y_i) + \lambda \Omega(\boldsymbol{\theta})$$

- Loss function $\ell(f_{\theta}(\mathbf{x}_i), y_i)$: cost of the model's output $f_{\theta}(\mathbf{x})$ when the true value is y.
 - The empirical risk is $\hat{R}_n(\theta) = \sum_{i=1}^n \ell(f_{\theta}(\mathbf{x}_i), y_i)$
- Regularizer $\Omega(\theta)$ & trade-off parameter $\lambda \geq 0$:
 - Background information about preferred solutions
 - Provides numerical stability (Tikhonov-Regularizer)
 - allows for tighter error bounds (PAC-Theory)

Solve

$$\underset{\boldsymbol{\theta}}{\operatorname{argmin}} \sum_{i=1}^{n} \ell(f_{\boldsymbol{\theta}}(\mathbf{x}_i), y_i) + \lambda \Omega(\boldsymbol{\theta})$$

Linear model:

$$\underset{\boldsymbol{\theta}}{\operatorname{argmin}} \sum_{i=1}^{n} \ell(\mathbf{x}_{i}^{\mathrm{T}}\boldsymbol{\theta}, y_{i}) + \lambda \Omega(\boldsymbol{\theta})$$

Linear model: solve

$$\underset{\boldsymbol{\theta}}{\operatorname{argmin}} \sum_{i=1}^{n} \ell(\mathbf{x}^{\mathrm{T}}\boldsymbol{\theta}, y_i) + \lambda \Omega(\boldsymbol{\theta})$$

- How to find solution:
 - Classification: No analytic solution but numeric solutions (gradient descent, cutting plane, interior point method)

Linear classification model: minimize

$$L(\boldsymbol{\theta}) = \sum_{i=1}^{n} \ell(\mathbf{x}^{\mathrm{T}}\boldsymbol{\theta}, y_i) + \lambda \Omega(\boldsymbol{\theta})$$

- Gradient:
 - Vector of the derivatives
 with respect to each individual
 parameter
 - Direction of the steepest increase of the function $L(\theta)$.

$$abla L(\mathbf{\theta}) = \begin{pmatrix} \frac{\partial L(\mathbf{\theta})}{\partial \theta_1} \\ \vdots \\ \frac{\partial L(\mathbf{\theta})}{\partial \theta_m} \end{pmatrix}$$

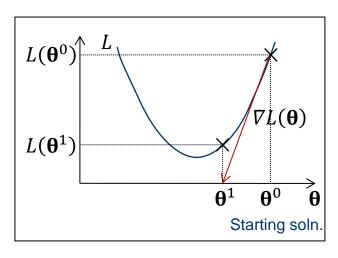
Linear classification model: minimize

$$L(\boldsymbol{\theta}) = \sum_{i=1}^{n} \ell(\mathbf{x}^{\mathrm{T}}\boldsymbol{\theta}, y_i) + \lambda \Omega(\boldsymbol{\theta})$$

Gradient descent method:

```
RegERM(Data: (\mathbf{x}_1, y_1), \dots, (\mathbf{x}_n, y_n))
Set \mathbf{\theta}^0 = \mathbf{0} and t = 0
DO

Compute gradient \nabla L(\mathbf{\theta}^t)
Compute step size \alpha^t
Set \mathbf{\theta}^{t+1} = \mathbf{\theta}^t - \alpha^t \nabla L(\mathbf{\theta}^t)
Set t = t+1
WHILE \|\mathbf{\theta}^t - \mathbf{\theta}^{t+1}\| > \varepsilon
RETURN \mathbf{\theta}^t
```



Linear classification model: minimize

$$L(\boldsymbol{\theta}) = \sum_{i=1}^{n} \ell(\mathbf{x}^{\mathrm{T}}\boldsymbol{\theta}, y_i) + \lambda \Omega(\boldsymbol{\theta})$$

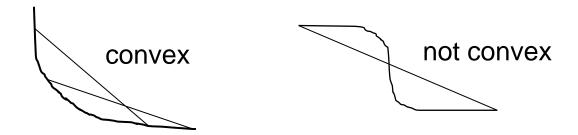
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Set t = t+1
WHILE \|\mathbf{\theta}^t - \mathbf{\theta}^{t+1}\| > \varepsilon
RETURN \mathbf{\theta}^t
```

- The step size α^t can be determined through
 - Line search
 - Barzilai-Borwein method
 - ...

- Properties of the gradient descent method:
 - Optimization criterion improved with every step.
 - Converges to the global minimum of the optimization criterion when this criterion is convex.
- The sum of convex functions is convex.
- Therefore, optimization criterion is convex if
 - Loss function is convex and
 - Regularizer is convex



ERM: Stochastic Gradient Method

- Idea: Determine the gradient for a random subset of the samples (e.g., a single instance).
- Less computation per optimization step, but only approximate descent direction.
- Optimization criterion with regularizer in sum:

$$L(\mathbf{\theta}) = \sum_{i=1}^{n} \left[\ell(f_{\mathbf{\theta}}(\mathbf{x}_i), y_i) + \frac{\lambda}{n} \Omega(\mathbf{\theta}) \right]$$

Stochastic gradient for a single instance:

$$\nabla_{\mathbf{x}_i} L(\mathbf{\theta}) = \frac{\partial}{\partial \mathbf{\theta}} \ell(f_{\mathbf{\theta}}(\mathbf{x}_i), y_i) + \frac{\lambda}{n} \frac{\partial}{\partial \mathbf{\theta}} \Omega(\mathbf{\theta})$$

ERM: Stochastic Gradient Method

Approximate gradient using single examples.

```
RegERM-Stoch (Data: (\mathbf{x}_1, y_1), \dots, (\mathbf{x}_n, y_n))

Set \mathbf{\theta}^0 = \mathbf{0} and t = 0

DO

Shuffle data randomly

FOR i = 1, \dots, n

Compute subset gradient \nabla_{\mathbf{x}_i} L(\mathbf{\theta}^t)

Compute step size \alpha^t

Set \mathbf{\theta}^{t+1} = \mathbf{\theta}^t - \alpha^t \nabla_{\mathbf{x}_i} L(\mathbf{\theta}^t)

Set t = t + 1

END

WHILE \|\mathbf{\theta}^t - \mathbf{\theta}^{t+1}\| > \varepsilon

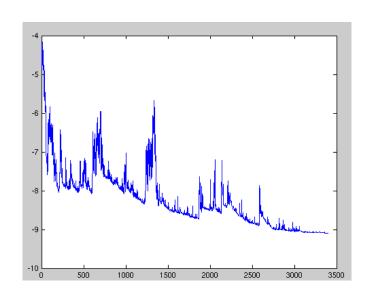
RETURN \mathbf{\theta}^t
```

ERM: Stochastic Gradient Method

- In every step only one summand of the optimization criterion is improved.
- The total optimization criterion can be worsened by these individual steps.
- Converges to the optimum if the step sizes satisfy:

$$\sum_{t=1}^{\infty} \alpha^t = \infty$$
 and $\sum_{t=1}^{\infty} (\alpha^t)^2 < \infty$

(Robbins & Monro, 1951)



ERM: Loss Functions for Classification

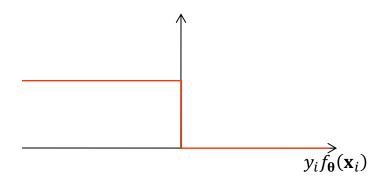
Zero-one loss:

$$\ell_{0/1}(f_{\boldsymbol{\theta}}(\mathbf{x}_i), y_i) = \begin{cases} 1 & -y_i f_{\boldsymbol{\theta}}(\mathbf{x}_i) \neq y_i \\ 0 & -y_i f_{\boldsymbol{\theta}}(\mathbf{x}_i) \leq 0 \end{cases}$$

$$\operatorname{sign}(f_{\boldsymbol{\theta}}(\mathbf{x}_i)) \neq y_i$$

$$\operatorname{sign}(f_{\boldsymbol{\theta}}(\mathbf{x}_i)) \leq 0$$

$$\operatorname{sign}(f_{\boldsymbol{\theta}}(\mathbf{x}_i)) = y_i$$



ERM: Loss Functions for Classification

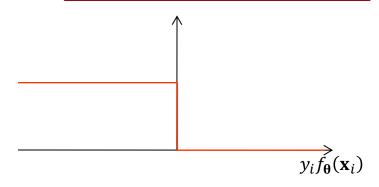
Zero-one loss:

$$\ell_{0/1}(f_{\theta}(\mathbf{x}_i), y_i) = \begin{cases} 1 & -y_i f_{\theta}(\mathbf{x}_i) \neq y_i \\ 0 & -y_i f_{\theta}(\mathbf{x}_i) \leq 0 \end{cases}$$

$$\operatorname{sign}(f_{\theta}(\mathbf{x}_i)) \neq y_i$$

$$\operatorname{sign}(f_{\theta}(\mathbf{x}_i)) \neq y_i$$

Zero-one loss is not convex ⇒ difficult to minimize!



ERM: Loss Functions for Classification

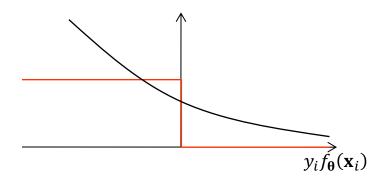
Zero-one loss:

$$\ell_{0/1}(f_{\boldsymbol{\theta}}(\mathbf{x}_i), y_i) = \begin{cases} 1 & -y_i f_{\boldsymbol{\theta}}(\mathbf{x}_i) \neq y_i \\ 0 & -y_i f_{\boldsymbol{\theta}}(\mathbf{x}_i) \leq 0 \end{cases}$$

$$\underset{\text{sign}(f_{\boldsymbol{\theta}}(\mathbf{x}_i)) = y_i}{\text{sign}(f_{\boldsymbol{\theta}}(\mathbf{x}_i)) = y_i}$$

Logistic loss:

$$\ell_{log}(f_{\theta}(\mathbf{x}_i), y_i) = \log(1 + e^{-y_i f_{\theta}(\mathbf{x}_i)})$$



ERM: Loss Functions for Classification

Zero-one loss:

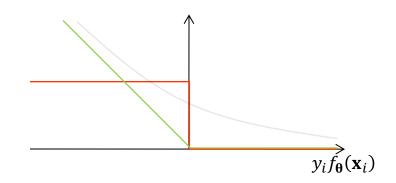
$$\ell_{0/1}(f_{\theta}(\mathbf{x}_i), y_i) = \begin{cases} 1 & -y_i f_{\theta}(\mathbf{x}_i) \neq y_i \\ 0 & -y_i f_{\theta}(\mathbf{x}_i) \leq 0 \end{cases}$$

$$\operatorname{sign}(f_{\theta}(\mathbf{x}_i)) \neq y_i$$

$$0 & \operatorname{sign}(f_{\theta}(\mathbf{x}_i)) = y_i$$



$$\ell_{log}(f_{\theta}(\mathbf{x}_i), y_i) = \log(1 + e^{-y_i f_{\theta}(\mathbf{x}_i)})$$



Perceptron loss:

$$\ell_p(f_{\theta}(\mathbf{x}_i), y_i) = \begin{cases} -y_i f_{\theta}(\mathbf{x}_i) & -y_i f_{\theta}(\mathbf{x}_i) > 0 \\ 0 & -y_i f_{\theta}(\mathbf{x}_i) \le 0 \end{cases} = \max(0, -y_i f_{\theta}(\mathbf{x}_i))$$

ERM: Loss Functions for Classification

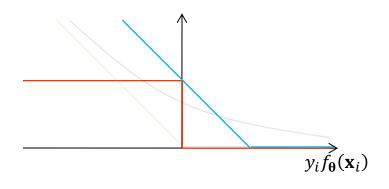
Zero-one loss:

Zero-one loss:
$$\sup_{\text{sign}(f_{\theta}(\mathbf{x}_i)) \neq y_i} \ell_{0/1}(f_{\theta}(\mathbf{x}_i), y_i) = \begin{cases} 1 & -y_i f_{\theta}(\mathbf{x}_i) \neq 0 \\ 0 & -y_i f_{\theta}(\mathbf{x}_i) \leq 0 \end{cases}$$

$$\sup_{\text{sign}(f_{\theta}(\mathbf{x}_i)) = y_i} \ell_{0}(\mathbf{x}_i) = y_i$$



$$\ell_{log}(f_{\theta}(\mathbf{x}_i), y_i) = \log(1 + e^{-y_i f_{\theta}(\mathbf{x}_i)})$$



Perceptron loss:

$$\ell_p(f_{\theta}(\mathbf{x}_i), y_i) = \begin{cases} -y_i f_{\theta}(\mathbf{x}_i) & -y_i f_{\theta}(\mathbf{x}_i) > 0 \\ 0 & -y_i f_{\theta}(\mathbf{x}_i) \le 0 \end{cases} = \max(0, -y_i f_{\theta}(\mathbf{x}_i))$$

Hinge loss:

$$\ell_h(f_{\theta}(\mathbf{x}_i), y_i) = \begin{cases} 1 - y_i f_{\theta}(\mathbf{x}_i) & 1 - y_i f_{\theta}(\mathbf{x}_i) > 0 \\ 0 & 1 - y_i f_{\theta}(\mathbf{x}_i) \le 0 \end{cases} = \max(0.1 - y_i f_{\theta}(\mathbf{x}_i))$$

ERM: Regularizers for Classification

- Idea: use as few attributes as possible:
 - $\Omega_0(\mathbf{\theta}) \propto \|\mathbf{\theta}\|_0 = \text{ number of } j \text{ with } \theta_j \neq 0$

 Ω_0 is not convex \Rightarrow difficult to minimize!

Manhattan norm (encourages scarcity):

$$\Omega_1(\mathbf{\theta}) \propto \|\mathbf{\theta}\|_1 = \sum_{i=1}^m |\theta_i|$$

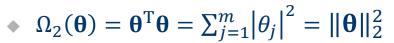
- Squared Euclidean norm (encourages small weights):
 - $\Omega_2(\boldsymbol{\theta}) \propto \|\boldsymbol{\theta}\|_2^2 = \sum_{j=1}^m \theta_j^2$

- Class $y \in \{-1, +1\}$
- Loss function:

$$\ell_h(f_{\theta}(\mathbf{x}_i), y_i) = \begin{cases} 1 - y_i f_{\theta}(\mathbf{x}_i) & \text{if } 1 - y_i f_{\theta}(\mathbf{x}_i) > 0 \\ 0 & \text{if } 1 - y_i f_{\theta}(\mathbf{x}_i) \le 0 \end{cases}$$

$$= \max(0, 1 - y_i f_{\theta}(\mathbf{x}_i))$$





Loss function is 0, if...

$$\begin{split} &\sum_{i=1}^{n} \max \left(0,1-y_{i}f_{\theta}(\mathbf{x}_{i})\right) = 0 \\ &\Leftrightarrow \ \forall_{i=1}^{n} \colon y_{i}f_{\theta}(\mathbf{x}_{i}) \geq 1 \\ &\Leftrightarrow \ \forall_{i=1}^{n} \colon y_{i}\mathbf{x}_{i}^{\mathsf{T}}\boldsymbol{\theta} \geq 1 \\ &\Leftrightarrow \ \forall_{i=1}^{n} \colon y_{i}\mathbf{x}_{i}^{\mathsf{T}}\frac{\boldsymbol{\theta}}{\|\boldsymbol{\theta}\|_{2}} \geq \frac{1}{\|\boldsymbol{\theta}\|_{2}} \\ &\Leftrightarrow \ \forall_{i=1}^{n} \colon x_{i}^{\mathsf{T}}\frac{\boldsymbol{\theta}}{\|\boldsymbol{\theta}\|_{2}} \left\{ \geq \frac{1}{\|\boldsymbol{\theta}\|_{2}} \quad \text{if } y_{i} = +1 \\ &\Leftrightarrow \ \forall_{i=1}^{n} \colon \mathbf{x}_{i}^{\mathsf{T}}\frac{\boldsymbol{\theta}}{\|\boldsymbol{\theta}\|_{2}} \left\{ \geq \frac{1}{\|\boldsymbol{\theta}\|_{2}} \quad \text{if } y_{i} = -1 \\ \end{split}$$

Loss function is 0, if...

$$\sum_{i=1}^{n} \max(0, 1 - y_i f_{\theta}(\mathbf{x}_i)) = 0$$

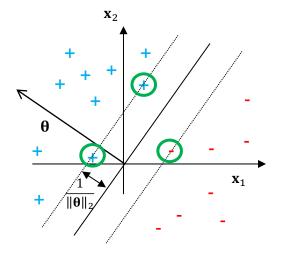
$$\Leftrightarrow \forall_{i=1}^{n} \colon y_i f_{\theta}(\mathbf{x}_i) \ge 1$$

$$\Leftrightarrow \forall_{i=1}^{n} \colon y_i \mathbf{x}_i^T \theta \ge 1$$

$$\Leftrightarrow \forall_{i=1}^{n} \colon y_i \mathbf{x}_i^T \frac{\theta}{\|\theta\|_2} \ge \frac{1}{\|\theta\|_2}$$

$$\Leftrightarrow \forall_{i=1}^{n} \colon \mathbf{x}_i^T \frac{\theta}{\|\theta\|_2}$$

$$\begin{cases} \ge \left(\frac{1}{\|\theta\|_2}\right) & \text{if } y_i = +1 \\ < \frac{1}{\|\theta\|_2} & \text{if } y_i = -1 \end{cases}$$

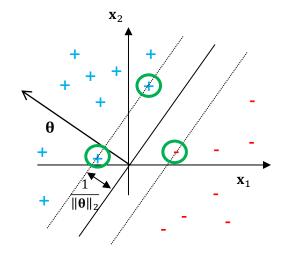


Distance of closest point to plane (margin)

For loss to be 0, all training samples musτ

- lie on the correct side of separating plane,
- and have a minimal margin (gap) of $\frac{1}{\|\mathbf{\theta}\|_2}$ to the plane.

- Loss function is 0 if all training samples have a margin of at least \(\frac{1}{\|\theta\|_2}\).
 - Points that lie $\frac{1}{\|\theta\|_2}$ from the plane are support vectors

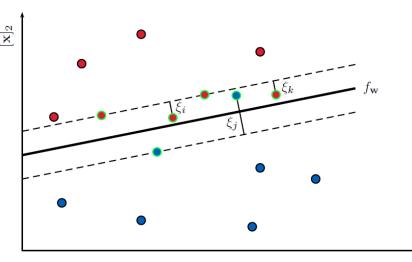


- Regularizer
 - $\Omega_2(\boldsymbol{\theta}) = \boldsymbol{\theta}^T \boldsymbol{\theta} = \|\boldsymbol{\theta}\|_2^2$; is zero only if $\boldsymbol{\theta} = \boldsymbol{0}$
 - Minimizing $\Omega_2(\mathbf{\theta}) \iff \text{maximizing margin } \frac{1}{\|\mathbf{\theta}\|_2}$
- SVM is also referred to as a *large margin classifier* because its optimization criterion is minimized by the plane with the largest margin from any sample.

- If loss function >0, some instances violate margin.
- Loss function as a sum of slack terms

 $\sum_{i=1}^{n} \max(0,1-y_i f_{\theta}(\mathbf{x}_i)) = \sum_{i=1}^{n} \xi_i$ $\xi_i = \max(0,1-y_i f_{\theta}(\mathbf{x}_i))$

Slack term or margin violation



 $[\mathbf{x}]_1$

Points with non-zero slack are support vectors

- Minimize hinge loss and L2-norm $\|\mathbf{\theta}\|_2^2 = \mathbf{\theta}^T \mathbf{\theta}$ of the parameter vector.
- Hinge loss is positive for a sample if the sample has a distance (margin) of less than $\frac{1}{\|\theta\|_2}$ to the separating hyperplane.
- SVM thereby finds the hyperplane with the greatest margin that separates the most possible samples. It trades off between
 - The size of the margin $\frac{1}{\|\theta\|_2}$
 - And the sum of the slack errors $\sum_{i=1}^{n} \xi_i$

Linear classification model: minimize

$$L(\boldsymbol{\theta}) = \sum_{i=1}^{n} \left[\max(0, 1 - y_i \mathbf{x}_i^{\mathrm{T}} \boldsymbol{\theta}) + \frac{\lambda}{n} \boldsymbol{\theta}^{\mathrm{T}} \boldsymbol{\theta} \right]$$

Gradient:

$$\nabla L(\boldsymbol{\theta}) = \sum_{i=1}^{n} \nabla_{\mathbf{x}_i} L(\boldsymbol{\theta})$$

Stochastic gradient for x_i:

$$\nabla_{\mathbf{x}_i} \mathbf{L}(\boldsymbol{\theta}) =$$

Linear classification model: minimize

$$L(\boldsymbol{\theta}) = \sum_{i=1}^{n} \left[\max(0, 1 - y_i \mathbf{x}_i^{\mathrm{T}} \boldsymbol{\theta}) + \frac{\lambda}{n} \boldsymbol{\theta}^{\mathrm{T}} \boldsymbol{\theta} \right]$$

Gradient:

$$\nabla L(\boldsymbol{\theta}) = \sum_{i=1}^{n} \nabla_{\mathbf{x}_i} L(\boldsymbol{\theta})$$

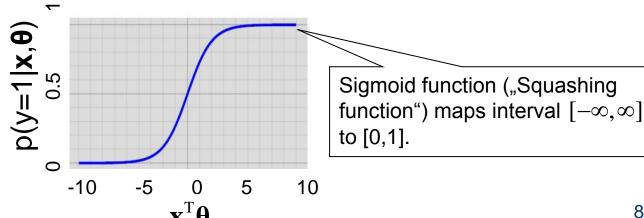
• Stochastic gradient for x_i :

$$\nabla_{\mathbf{x}_i} \boldsymbol{L}(\boldsymbol{\theta}) = \begin{cases} \frac{2\lambda}{n} \boldsymbol{\theta} & \text{if } y_i \mathbf{x}_i^{\mathrm{T}} \boldsymbol{\theta} > 1 \\ \frac{2\lambda}{n} \boldsymbol{\theta} - y_i \mathbf{x}_i & \text{if } y_i \mathbf{x}_i^{\mathrm{T}} \boldsymbol{\theta} < 1 \end{cases}$$

- $L(\theta)$ can be minimized using stochastic gradient descent method ("Pegasos")
 - Very fast, often used in practice
- $L(\theta)$ can be minimized using gradient descent method ("Primal SVM")

Logistic Regression

- "Logistic regression" is a model for classification!
- For now, binary classification with $y_i \in \{-1,1\}$.
- Need: model for $p(y | \mathbf{x}, \boldsymbol{\theta})$
 - Model defines probability $p(y = 1 | \mathbf{x}, \mathbf{\theta})$.
 - Probability $p(y = -1 \mid \mathbf{x}, \mathbf{\theta}) = 1 p(y = 1 \mid \mathbf{x}, \mathbf{\theta}).$
- Idea: transformation of a linear model $\mathbf{x}^{\mathrm{T}}\mathbf{\theta}$.



Logistic Regression

- Model logistic regression
 - Given by parameter vector $\mathbf{\theta} \in \square^m$.
 - Defines conditional distribution $p(y | \mathbf{x}, \mathbf{\theta})$ by

$$p(y=1 | \mathbf{x}, \mathbf{\theta}) = \sigma(\mathbf{x}^{\mathrm{T}}\mathbf{\theta}) = \frac{1}{1 + \exp(-\mathbf{x}^{\mathrm{T}}\mathbf{\theta})}$$

$$\sigma(z)$$

$$p(y = -1 | \mathbf{x}, \mathbf{\theta}) = 1 - p(y = 1 | \mathbf{x}, \mathbf{\theta})$$

■ Prediction function $f_{\theta}: \square^m \to \{0,1\}$:

$$f_{\mathbf{\theta}}(\mathbf{x}) = \begin{cases} 1: \ \sigma(\mathbf{x}^{\mathrm{T}}\mathbf{\theta}) \ge 0.5 \\ 0: \ \text{sonst} \end{cases}$$

Learning Logistic Regression Models

MAP model: minimize regularized loss.

$$\begin{aligned} \mathbf{\theta}_{\text{MAP}} &= \arg \max_{\mathbf{\theta}} P(\mathbf{y} \mid \mathbf{X}, \mathbf{\theta}) p(\mathbf{\theta}) \\ &= \arg \min_{\mathbf{\theta}} \sum_{i=1}^{n} \log \left(1 + \exp(-y_{i} \mathbf{x}_{i}^{\text{T}} \mathbf{\theta}) \right) + \frac{1}{2\sigma_{p}^{2}} |\mathbf{\theta}|^{2} \\ &\qquad \qquad \qquad \text{loss function} \end{aligned}$$

- Convex optimization problem, global minimum.
- Compare earlier lecture on "Linear models".

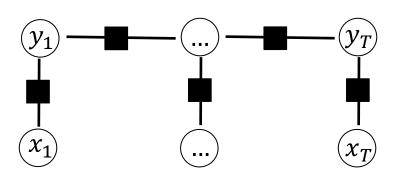
Overview

- Graphical models.
- The n-gram model.
- Hidden Markov model.
- Linear classification models.
- Conditional random fields.
- PCFGs
- Forward- and backpropagation in neural networks.
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- LSTM networks.

Factorization:

$$P(\mathbf{x}, \mathbf{y} | \mathbf{\theta}) = \frac{1}{Z} \prod_{t=1}^{T} \exp\{\psi(y_t, y_{t+1}) + \psi(y_t, x_t)\}$$

$$= \frac{1}{Z} \prod_{t=1}^{T} \exp\left\{\sum_{i,j} \theta_{ij} [y_t = i, y_{t+1} = j] + \sum_{i,o} \theta_{io} [y_t = i, x_t = o]\right\}$$



Factorization:

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$$= \frac{1}{Z} \exp\{\mathbf{\theta}^T \Phi(\mathbf{x}, \mathbf{y})\}$$

$$y_1 \longrightarrow \mathbf{w} \longrightarrow \mathbf{y}_T$$

Conditional random field:

$$P(\mathbf{y}|\mathbf{x}, \mathbf{\theta}) = \frac{P(\mathbf{x}, \mathbf{y}|\mathbf{\theta})}{\sum_{\mathbf{y}'} P(\mathbf{x}, \mathbf{y}'|\mathbf{\theta})}$$

$$= \frac{1}{Z} \prod_{t=1}^{T} \exp\{\psi(y_t, y_{t+1}) + \psi(y_t, x_t)\}$$

$$= \frac{1}{Z} \prod_{t=1}^{T} \exp\left\{\sum_{i,j} \theta_{ij} [y_t = i, y_{t+1} = j] + \sum_{i,o} [y_t = i, x_t = o]\right\}$$

$$= \frac{\exp\{\mathbf{\theta}^T \Phi(\mathbf{x}, \mathbf{y})\}}{\sum_{\mathbf{y}'} \exp\{\mathbf{\theta}^T \Phi(\mathbf{x}, \mathbf{y}')\}}$$

Factorization:

$$P(\mathbf{x}, \mathbf{y} | \mathbf{\theta}) = \frac{1}{Z} \exp{\{\mathbf{\theta}^{\mathrm{T}} \Phi(\mathbf{x}, \mathbf{y})\}}$$

With

$$\Phi(\mathbf{x}, \mathbf{y}) = \begin{pmatrix} \sum_{t} [y_{t} = 1, y_{t+1} = 1] \\ \vdots \\ \sum_{t} [y_{t} = N, y_{t+1} = N] \\ \sum_{t} [x_{t} = o_{1}, y_{t} = 1] \\ \vdots \\ \sum_{t} [x_{t} = o_{M}, y_{t} = N] \end{pmatrix}$$

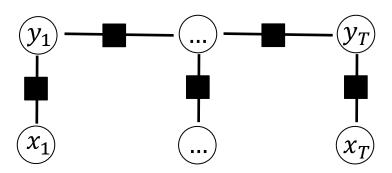
$$y_{1} \qquad y_{T}$$

Decision function:

$$f_{\mathbf{\theta}}(\mathbf{x}, \mathbf{y}) = \mathbf{\theta}^{\mathrm{T}} \Phi(\mathbf{x}, \mathbf{y})$$

With

$$\Phi(\mathbf{x}, \mathbf{y}) = \begin{pmatrix} \sum_{t} [y_{t} = 1, y_{t+1} = 1] \\ \vdots \\ \sum_{t} [y_{t} = N, y_{t+1} = N] \\ \sum_{t} [x_{t} = o_{1}, y_{t} = 1] \\ \vdots \\ \sum_{t} [x_{t} = o_{M}, y_{t} = N] \end{pmatrix}$$



Undirected Sequence Models: Prediction

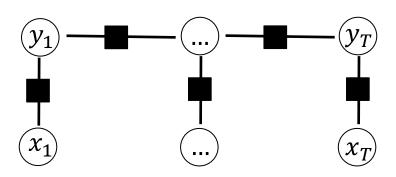
Conditional random field: most likely output

$$\hat{\mathbf{y}} = \arg \max_{\mathbf{y}} P(\mathbf{y}|\mathbf{x}, \mathbf{\theta}) = \arg \max_{\mathbf{y}} \frac{1}{Z} \exp\{\mathbf{\theta}^{T} \Phi(\mathbf{x}, \mathbf{y})\}\$$

$$= \arg \max_{\mathbf{y}} \mathbf{\theta}^{T} \Phi(\mathbf{x}, \mathbf{y})$$

Decision function:

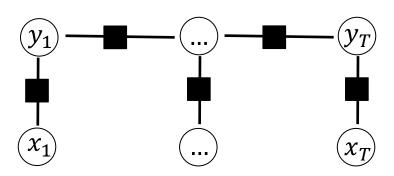
$$\hat{\mathbf{y}} = \underset{\mathbf{y}}{\operatorname{argmax}} f_{\mathbf{\theta}}(\mathbf{x}, \mathbf{y}) = \mathbf{\theta}^{\mathrm{T}} \Phi(\mathbf{x}, \mathbf{y})$$



Undirected Sequence Models: Prediction

$$\hat{\mathbf{y}} = \underset{\mathbf{y}}{\operatorname{argmax}} f_{\mathbf{\theta}}(\mathbf{x}, \mathbf{y}) = \mathbf{\theta}^{\mathrm{T}} \Phi(\mathbf{x}, \mathbf{y})$$

- Maximization with variant of Viterbi allgorithm.
 - Scores instead of log-probabilities

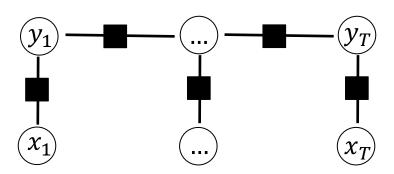


Conditional Random Fields: Inference

Conditional random field: probability

$$P(\mathbf{y}|\mathbf{x}, \mathbf{\theta}) = \frac{\exp{\{\mathbf{\theta}^{\mathrm{T}} \Phi(\mathbf{x}, \mathbf{y})\}}}{\sum_{\mathbf{y}'} \exp{\{\mathbf{\theta}^{\mathrm{T}} \Phi(\mathbf{x}, \mathbf{y}')\}}}$$

■ Denominator $\sum_{\mathbf{y}'} \exp\{\mathbf{\theta}^T \Phi(\mathbf{x}, \mathbf{y}')\}$ inferred by variant of the Viterbi algorithm.



Conditional Random Field: Learning

Optimization criterion

$$\arg \max_{\mathbf{\theta}} \prod_{i=1}^{n} P(\mathbf{y}_{i} | \mathbf{x}_{i}, \mathbf{\theta}) P(\mathbf{\theta})$$

$$= \arg \max_{\mathbf{\theta}} \sum_{i=1}^{n} \mathbf{\theta}^{T} \Phi(\mathbf{x}_{i}, \mathbf{y}_{i}) + \Omega(\mathbf{\theta})$$

- Optimization for instance by stochastic gradient descent.
- Requires labeled training data (y_i visible).

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PCFG: Definitions

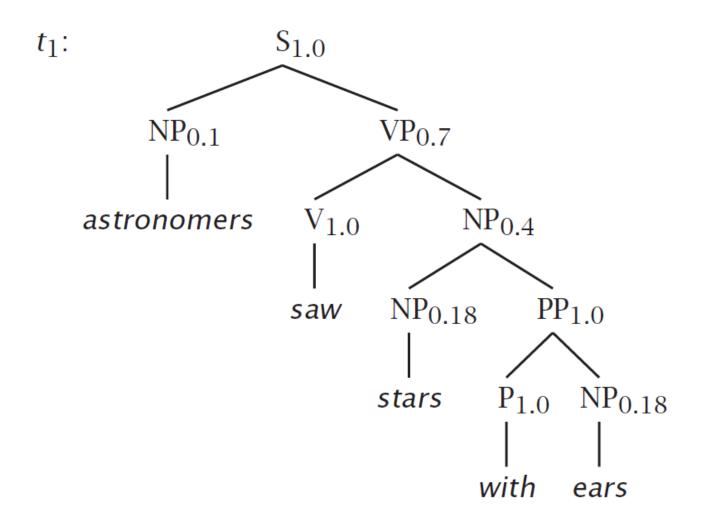
- Terminal symbols: $\{w^k\}$, $k = 1 \dots v$.
- Nonterminal symbols: $\{N^i\}$, $i = 1 \dots n$.
- Starting symbol N^1 .
- Productions: $\{N^i \to \xi^j\}$, whete ξ^j is a sequence of terminals and nonterminals.
- Probabilities $P(N^i \to \xi^j)$, such that for all i: $\sum_j P(N^i \to \xi^j) = 1$.

PCFG: Definitions

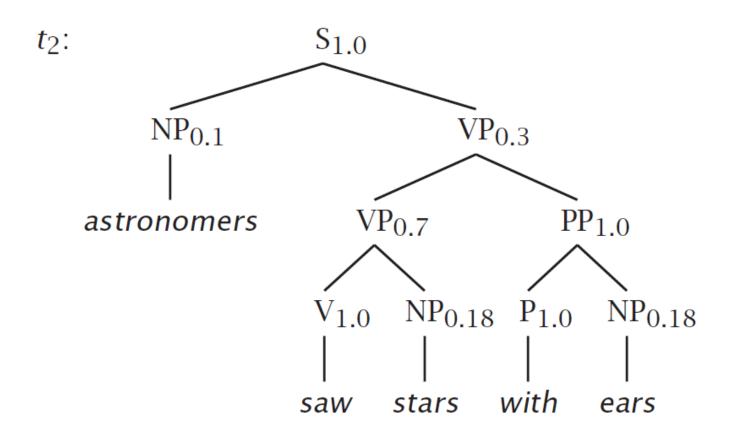
- Let $w_s, ..., w_e$ be a sentence.
- $N^i \Rightarrow^* w_S, ..., w_e$ if $w_S, ..., w_e$ can be derived from N^i .
- N_{Se}^{i} : derivation of w_{S} , ..., w_{e} from N^{i} .
- Example PCFG:

$S \rightarrow NP VP$	1.0	$NP \rightarrow NP PP$	0.4
$PP \rightarrow P NP$	1.0	NP → astronomers	0.1
$VP \rightarrow V NP$	0.7	NP → ears	0.18
$VP \rightarrow VP PP$	0.3	NP → saw	0.04
$P \rightarrow with$	1.0	NP → stars	0.18
V → saw	1.0	NP → telescopes	0.1

PCFG: Example Parse



PCFG: Example Parse



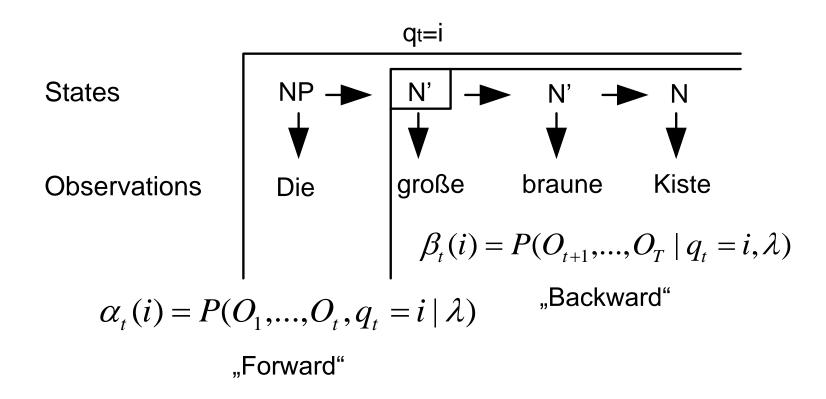
PCFG: Inference

- PCFGs answer three questions (much like HMMs):
 - What is the likelihood of a sentence given a PCFG?
 - What is the most likely parse tree?
 - What are the most likely PCFG parameters given a corpus of sentences and parse trees.

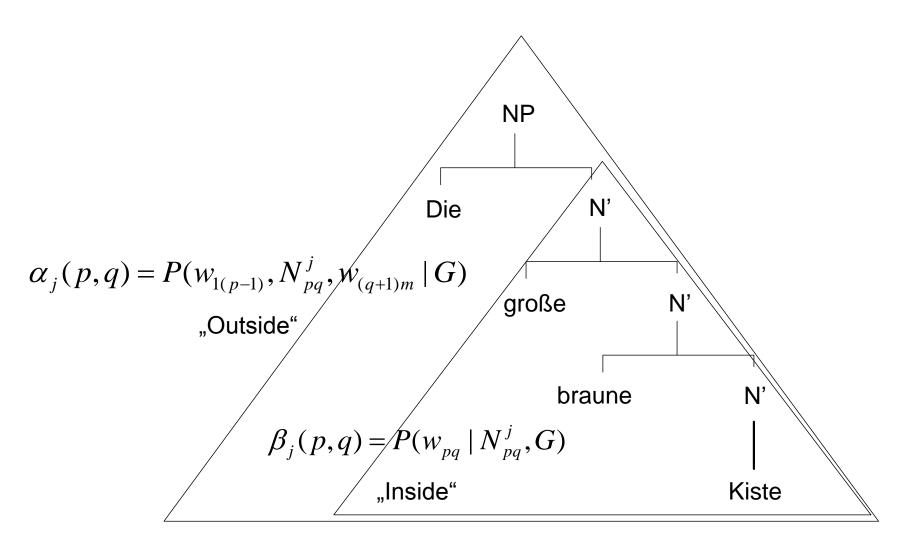
Chomsky Normal Form

- For each context-free grammar, there is a contextfree grammar in Chomsky normal form.
- Chomsky normal form has two types of productions:
 - $N^i \rightarrow N^j N^k$
 - $N^i \rightarrow w^j$
- Chomsky normal form allows a simplified treatment.

Forward / Backward – Inside / Outside



Forward / Backward – Inside / Outside



Inside / Outside

Outside probability:

$$\alpha_{j}(p,q) = P(w_{1(p-1)}, N_{pq}^{j}, w_{(q+1)m} \mid G)$$

Inside probability

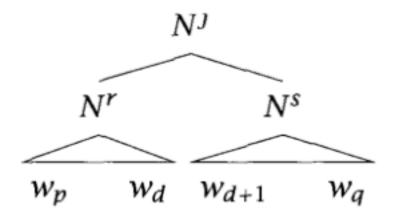
$$\beta_j(p,q) = P(w_{pq} \mid N_{pq}^j, G)$$

Likelihood of a chain of terms:

$$P(w_{1T} \mid G) = P(N^1 \Rightarrow^* w_{1T}, G)$$
$$= P(w_{1T} \mid N_{1T}^1, G) = \beta_1(1, T)$$

Calculating the Inside Probability

$$\beta_j(p,q) = P(w_{pq} \mid N_{pq}^j, G)$$



Calculating the Inside Probability

Base case:
$$\beta_j(k,k) = P(w_k \mid N_{kk}^j, G)$$

= $P(N_i \rightarrow w_k \mid G)$

Inductive step:

$$\begin{split} \beta_{j}(p,q) &= P(w_{pq}|N_{pq}^{j},G) \\ &= \sum_{r,s} \sum_{d=p}^{q-1} P(w_{pd},N_{pd}^{r},w_{(d+1)q},N_{(d+1)q}^{s}|N_{pq}^{j},G) \\ &= \sum_{r,s} \sum_{d=p}^{q-1} P(N_{pd}^{r},N_{(d+1)q}^{s}|N_{pq}^{j},G)P(w_{pd}|N_{pq}^{j},N_{pd}^{r},N_{(d+1)q}^{s},G) \\ &\times P(w_{(d+1)q}|N_{pq}^{j},N_{pd}^{r},N_{(d+1)q}^{s},w_{pd},G) \\ &= \sum_{r,s} \sum_{d=p}^{q-1} P(N_{pd}^{r},N_{(d+1)q}^{s}|N_{pq}^{j},G)P(w_{pd}|N_{pd}^{r},G) \\ &\times P(w_{(d+1)q}|N_{(d+1)q}^{s},G) \\ &\times P(w_{(d+1)q}|N_{(d+1)q}^{s},G) \end{split}$$

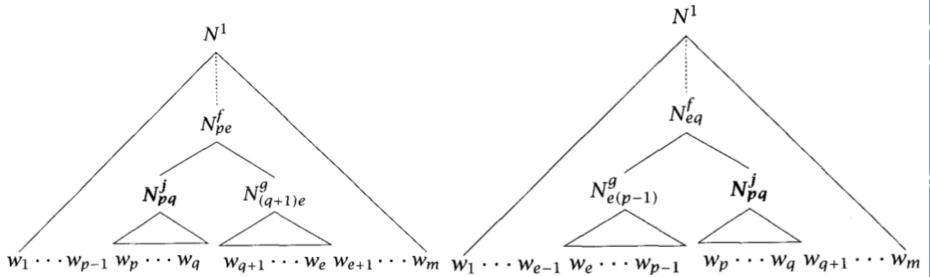
Example: Calculating Inside Probability

$S \rightarrow NP VP$	1.0	$NP \rightarrow NP PP$	0.4
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$P \rightarrow with$	1.0	NP → stars	0.18
V → saw	1.0	NP → telescopes	0.1

	1	2	3	4	5
1	$\beta_{NP} = 0.1$		$\beta_{S} = 0.0126$		$\beta_{S} = 0.0015876$
2		$\beta_{NP} = 0.04$	$\beta_{VP} = 0.126$		$\beta_{VP} = 0.015876$
		$\beta_{V} = 1.0$			
3			$\beta_{NP} = 0.18$		$\beta_{NP} = 0.01296$
4				$\beta_{P} = 1.0$	$\beta_{PP} = 0.18$
5					$\beta_{NP} = 0.18$
	astronomers	saw	stars	with	ears

Calculating the Outside Probability

$$\alpha_{j}(p,q) = P(w_{1(p-1)}, N_{pq}^{j}, w_{(q+1)m} \mid G)$$



Calculating the Outside Probability

- Base case: $\alpha_1(1,T) = 1$; $\alpha_j(1,T) = 0$ für $j \ne 1$
- Inductive step:

$$\begin{aligned} &\alpha_{j}(p,q) = P(w_{1(p-1)}, N_{pq}^{j}, w_{(q+1)m} \mid G) \\ &= \left(\sum_{f,g} \sum_{e=q+1}^{m} P(w_{1(p-1)}, w_{(q+1)m}, N_{pe}^{f}, N_{pq}^{j}, N_{(q+1)e}^{g}) \right) \\ &+ \left(\sum_{f,g} \sum_{e=1}^{p-1} P(w_{1(p-1)}, w_{(q+1)m}, N_{eq}^{f}, N_{e(p-1)}^{g}, N_{pq}^{g}) \right) \\ &= \left(\sum_{f,g} \sum_{e=q+1}^{m} P(w_{1(p-1)}, w_{(q+1)m}, N_{pe}^{f}) P(N_{pq}^{j} N_{(q+1)e}^{g} \mid N_{pe}^{f}) P(w_{(q+1)e} \mid N_{(q+1)e}^{g}) \right) \\ &+ \left(\sum_{f,g} \sum_{e=q+1}^{p-1} P(w_{1(e-1)}, w_{(q+1)m}, N_{eq}^{f}) P(N_{e(p-1)}^{g}, N_{pq}^{j} \mid N_{eq}^{f}) P(w_{e(p-q)} \mid N_{e(p-1)}^{g}) \right) \\ &= \left(\sum_{f,g} \sum_{e=q+1}^{m} \alpha_{f}(p, e) P(N^{f} \rightarrow N^{j} N^{g}) \beta_{g}(q+1, e) \right) \\ &+ \left(\sum_{f,g} \sum_{e=q+1}^{p-1} \alpha_{f}(e, q) P(N^{f} \rightarrow N^{g} N^{j}) \beta_{g}(e, p-1) \right) \end{aligned}$$

Most Likely Parse Tree

- Viterbi algorithm for PCFG
- $O(m^3n^3)$
- $\delta_i(p,q)$: maximum inside probability for a parse of subtree N_{pq}^i .
- Initialization: $\delta_i(p,p) = P(N^i \to w_p)$
- Induction: $\delta_i(p,q) = \max_{j,k,p \le r < q} P(N^i \to N^j N^k) \delta_j(p,r) \delta_k(r+1,q)$
- Save best parse tree:

$$\psi_i(p,q) = \operatorname{arg\,max}_{(j,k,r)} P(N^i \to N^j N^k) \delta_j(p,r) \delta_k(r+1,q)$$

Reconstruct parse tree: if $\psi_i(p,q) = (j,k,r)$, then left branch starts with N_{pr}^j , right branch with $N_{(r+1)q}^k$

PCFG: Parameter Estimation

- Grammar has to be known; only probabilities $P(N^i \to N^j N^k)$ are estimated.
- If corpus is annotated with parse trees: count how frequently each rule is used + Laplace smoothing.
- Without parse tree annotation: use EM algorithm (Baum Welch for PCFGs).

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